

Generalized birthday problem for October 12

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- 2 Theorem I with examples
- 3 Theorems II and III
- 4 Proof overview of Theorem I
- 5 Conclusion

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- The answer to this question is very well known.
- In this case the number of student pairs with a common birthday is approximate Poisson with mean $\lambda = \frac{\binom{n}{2}}{c}$.
- Consequently the chance of at least one common birthday is approximately $1 - e^{-\lambda}$.

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- Here $|E(G)|$ is the number of edges in G , i.e. the total number of friendship pairs in the class.
- The classical birthday problem is a special case with $G = K_n$, where everyone knows everyone else.

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- In this case things are more interesting, and the answer depends on the graph in a more delicate way (than just the number of edges).

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- What is the limiting distribution of T_n , the number of monochromatic edges of color 1?
- Note that

$$T_n = \sum_{i < j} G_n(i, j) X_i X_j,$$

where (X_1, \dots, X_n) are i.i.d. $\text{Bern}(p_n)$, with $p_n = \frac{1}{c_n}$.

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- We will assume that the sequence $\{G_n, p_n\}_{n \geq 1}$ are chosen such that $\mathbb{E}T_n = \frac{|E(G_n)|}{c_n^2} = O(1)$.

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- Thus $T_n \xrightarrow{D} \binom{S}{2}$, where $S \sim \text{Pois}(1)$.

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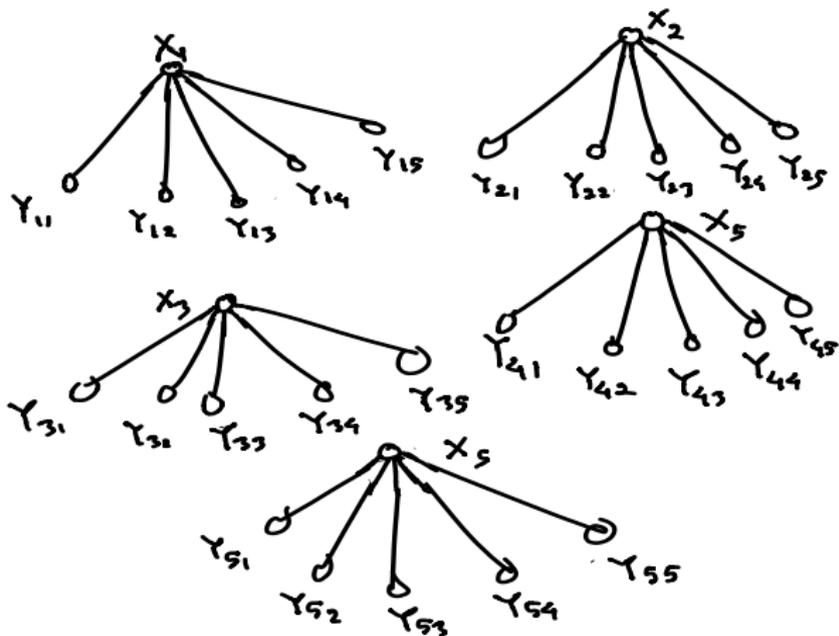
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EG : $n = 25$, 5 COPIES OF $K_{1,5}$

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$$T_n = \sum_{i=1}^{\sqrt{n}} X_i \sum_{j=1}^{\sqrt{n}} Y_{ij},$$

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- This gives $T_n \xrightarrow{D} \text{Pois}(S)$, where $S \sim \text{Pois}(1)$.

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- As already seen, this class contains mixtures of Poissons, and Binomials of quadratic functions of Poissons.
- Also, can we characterize when is this limit exactly a Poisson?

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- The disjoint star example captures the contribution of edges between high degree vertices and low degree vertices.
- Finally, edges between low degree vertices gives rise to a Poisson limit.
- Using this philosophy, we partition the edge set into 3 types,

High \leftrightarrow High, High \leftrightarrow Low, Low \leftrightarrow Low.

Adjacency matrix \mapsto function on positive reals

- Define a function $W_{G_n}(\cdot, \cdot) : [0, \infty)^2 \mapsto [0, 1]$ by setting

$$\begin{aligned} W_{G_n}(x, y) &= 1 \text{ if } (\lceil xc_n \rceil, \lceil yc_n \rceil) \in E(G_n) \\ &= 0 \text{ otherwise} \end{aligned}$$

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- Let $d_{G_n}(x) := \int_0^\infty W_{G_n}(x, y) dy$ be the degree function.

First assumption (A1)

- Given two bounded measurable functions f, g from $[0, K]^2 \mapsto [0, 1]$, define the strong cut distance between f and g by

$$\sup_{A, B \subset [0, 1]} \left| \int_{A \times B} f(x, y) dx dy - \int_{A \times B} g(x, y) dx dy \right|.$$

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- In some sense W captures the limit of the dense part of the graph. Note that this assumption, along with Fatou's lemma automatically implies

$$\int_{[0, \infty)^2} |W(x, y)| dx dy < \infty.$$

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- In some sense $\Delta(x)$ counts the edges between the high and low degree vertices.

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- Assumption: (A3)

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- *The joint Mgf of (Q_1, Q_2) appears on the next slide.*

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- Q_1 arises from the edges between the high degree vertices, i.e. the dense part of the graph.

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- Then $I_2(f)$ is well defined, and

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- Suppose $p_n = \frac{1}{n}$, and G_n is a sequence of random graphs such that

$$\begin{aligned}\mathbb{P}(G_n(i, j) = 1) &= a_{11} \text{ if } i, j < \frac{n}{2} \\ &= a_{12} \text{ if } i < \frac{n}{2}, j \geq \frac{n}{2} \text{ or } i \geq \frac{n}{2}, j < \frac{n}{2}, \\ &= a_{22} \text{ if } i, j \geq \frac{n}{2}.\end{aligned}$$

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- Similar results apply to unequal blocks, or more than 2 blocks.

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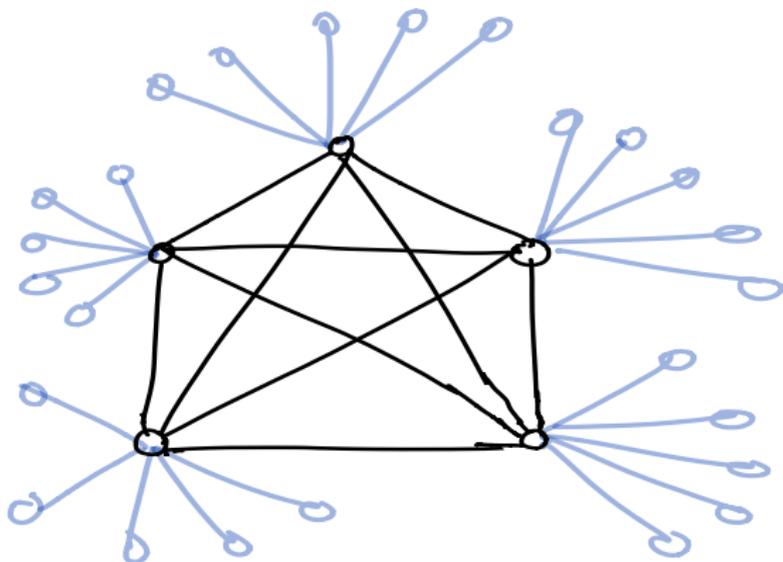
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- Then the entire graph G_n has $n + n^2 \sim n^2$ vertices, and $\binom{n}{2} + n^2 \sim \frac{3n^2}{2}$ edges.

Co-existence example



EG: $n = 5$

$$\text{TOTAL VERTICES} = 5 + 5 \times 5 = 30$$

$$\text{TOTAL EDGES} = \binom{5}{2} + 5 \times 5 = 35$$

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- This gives

$$Q_1 + Q_2 \xrightarrow{D} \binom{S}{2} + Pois(S), \text{ where } S \sim Pois(1).$$

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- This ensures that G_n has $\Theta(n^2)$ vertices, and $\Theta_P(n^2)$ many edges.
- For the choice $p_n = \frac{1}{n}$, our theorem applies with

$$W(x, y) = c_k \text{ if } x, y \in (r_{k-1}, r_k] \text{ for some } k \geq 1, \\ = 0 \text{ otherwise .}$$

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- Using our theorem gives

$$T_n \xrightarrow{D} \sum_{k=1}^{\infty} \text{Bin} \left(\binom{S_k}{2}, c_k \right) + \text{Pois} \left(\sum_{k=1}^{\infty} a_k S_k \right),$$

where $S_k \sim \text{Pois}(b_k)$ are mutually independent.

- 1 Introduction
- 2 Theorem I with examples
- 3 Theorems II and III
- 4 Proof overview of Theorem I
- 5 Conclusion

Theorem II

- Our first result shows that under (A1), (A2), (A3), the limit of T_n can be expressed as $\psi(W, d, \lambda)$ for suitable W, d, λ .

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- We claim that the class $\psi(W, d, \lambda)$ is closed under weak topology.

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Theorem (Bhattacharya-Mukherjee-M., AAP-2020)

If $\mathbb{E}T_n = O(1)$, the following are equivalent:

(i) $T_n \xrightarrow{D} \text{Pois}(\lambda)$.

(ii)

$$\lim_{M \rightarrow \infty} \lim_{n \rightarrow \infty} \mathbb{E}(T_{n,M}) = \lambda, \quad \lim_{M \rightarrow \infty} \lim_{n \rightarrow \infty} \text{Var}(T_{n,M}) = \lambda.$$

Theorem III

- Here

$$T_{n,M} := \sum_{i < j} G_n(i, j) X_i X_j 1\{d_i \leq M c_n, d_j \leq M c_n\}$$

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If $\mathbb{E}T_n \rightarrow \lambda$ and $\text{Var}(T_n) \rightarrow \lambda$, then $T_n \xrightarrow{D} \text{Pois}(\lambda)$.

- Compare this with the more well studied fourth moment phenomenon for the Gaussian distribution (see [Ivan Nourdin's](#) webpage for a list of papers on this topic).

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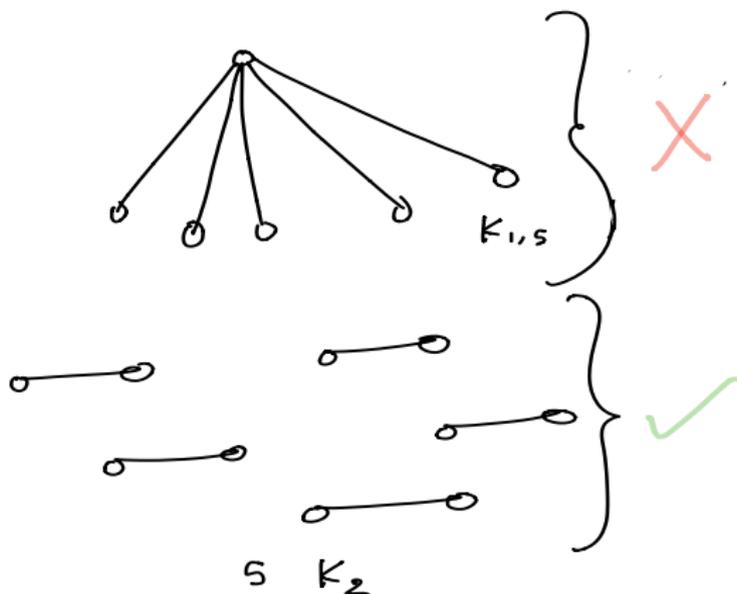
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- Similar results hold for sparse block models, and random regular graphs.

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- The truncated second moment result captures this behavior automatically.

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- By using a first moment computation using Markov's inequality, we show that the super high degree vertices do not contribute for M large.

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- We argue using method of moments that $T_{3,n}$ is asymptotically independent from $T_{1,n}$ and $T_{2,n}$.

Proof idea of Theorem-I

- Using
 - (i) (A1): strong cut metric convergence of W_{G_n} on $[0, K]^2$
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 - (ii) (A2): the convergence of the degree function d_{G_n} in $L^1[0, K]$,we argue that the joint moments of $T_{1,n}$ and $T_{2,n}$ converge.

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- We also show that the limiting moments determine their joint distribution.
- To identify the distribution of (Q_1, Q_2) , we compute the Mgf along a well chosen sequence of inhomogeneous random graphs.
- This gives the joint Mgf of (Q_1, Q_2) , thereby proving Theorem I.

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- We characterize the class of all possible limits of Bernoulli quadratic forms.
- As an application, we characterize exactly when is the limit a Poisson random variable.
- We apply our theorem to several examples, which includes both deterministic and random graphs.

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- Finally, our quadratic form is (in terms of) the adjacency matrix of a simple graph. Does a similar analysis apply for **general** quadratic forms?



The End